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A Unifying Tutorial on Approximate Message Passing

Oliver Y. Feng

University of Cambridge UK o.feng@statslab.cam.ac.uk

Ramji Venkataramanan

University of Cambridge UK ramji.v@eng.cam.ac.uk

Cynthia Rush

Columbia University USA cgr2130@columbia.edu

Richard J. Samworth

University of Cambridge UK r.samworth@statslab.cam.ac.uk



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A Unifying Tutorial on Approximate Message Passing

Oliver Y. Feng¹, Ramji Venkataramanan², Cynthia Rush³ and Richard J. Samworth¹

¹Statistical Laboratory, University of Cambridge, UK;
o.feng@statslab.cam.ac.uk; r.samworth@statslab.cam.ac.uk
²Department of Engineering, University of Cambridge, UK;
ramji.v@eng.cam.ac.uk
³Department of Statistics, Columbia University, USA;
cgr2130@columbia.edu

ABSTRACT

Over the last decade or so, Approximate Message Passing (AMP) algorithms have become extremely popular in various structured high-dimensional statistical problems. Although the origins of these techniques can be traced back to notions of belief propagation in the statistical physics literature, our goals in this work are to present the main ideas of AMP from a statistical perspective and to illustrate the power and flexibility of the AMP framework. Along the way, we strengthen and unify many of the results in the existing literature.

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1

Introduction

Approximate Message Passing (AMP) refers to a class of iterative algorithms that have been successfully applied to a number of statistical estimation tasks such as linear regression (Bayati and Montanari, 2011; Donoho et al., 2009; Krzakala et al., 2012), generalised linear models (Mondelli and Venkataramanan, 2020; Rangan, 2011; Schniter and Rangan, 2014) and low-rank matrix estimation (Deshpande and Montanari, 2014; Deshpande et al., 2016; Kabashima et al., 2016; Lesieur et al., 2017; Matsushita and Tanaka, 2013; Montanari and Richard, 2016; Montanari and Venkataramanan, 2021; Rangan and Fletcher, 2018). Moreover, these techniques are also popular and practical in a variety of engineering and computer science applications such as imaging (Fletcher and Rangan, 2014; Metzler et al., 2017; Vila et al., 2015), communications (Barbier and Krzakala, 2017; Jeon et al., 2015; Rush et al., 2017; Schniter, 2011) and machine learning (El Alaoui *et al.*, 2018; Emami et al., 2020; Manoel et al., 2017; Pandit et al., 2020; Yang, 2019). AMP algorithms have two features that make them particularly attractive. First, they can easily be tailored to take advantage of prior information on the structure of the signal, such as sparsity or other constraints. Second, under suitable assumptions on a design or data matrix, AMP

theory provides precise asymptotic guarantees for statistical procedures in the high-dimensional regime where the ratio of the number of observations n to dimensions p converges to a constant (Bayati and Montanari, 2012; Donoho and Montanari, 2016; Donoho *et al.*, 2013; Sur *et al.*, 2017). More generally, AMP has been used to obtain lower bounds on the estimation error of first-order methods (Celentano *et al.*, 2020). In generalised linear models, low-rank matrix estimation and neural network models, it also plays a fundamental role in understanding the performance gap between information-theoretically optimal and computationally feasible estimators (Aubin *et al.*, 2019, 2020; Barbier *et al.*, 2019; Lelarge and Miolane, 2019; Reeves and Pfister, 2019). In these settings, it is conjectured that AMP achieves the optimal asymptotic estimation error among all polynomial-time algorithms (cf. Celentano and Montanari, 2022).

The purpose of this tutorial is to give a comprehensive and rigorous introduction to what AMP can offer, as well as to unify and formalise the core concepts within the large body of recent work in the area. We mention here that many of the original ideas of AMP were developed in the physics and engineering literature, and involved notions such as "loopy belief propagation" (e.g., Koller and Friedman, 2009, Section 11.3) and the "replica method" (e.g., Guo and Verdú, 2005; Krzakala et al., 2012; Mézard and Montanari, 2009; Tanaka, 2002; Rangan et al., 2009). Our starting point, however, will be an abstract AMP recursion, whose form depends on whether or not the data matrix is symmetric; we will study the symmetric case in detail, and then present the asymmetric version, which can be handled via a reduction argument. The striking and crucial feature of this recursion is that when the dimension is large, the empirical distribution of the coordinates of each iterate is approximately Gaussian, with limiting variance given by a scalar iteration called "state evolution".

Rigorous formulations of the key AMP property are given in Theorems 2.1 and 2.3 (for the symmetric case) and Theorem 2.5 (for the asymmetric case), which can be found in Sections 2.1 and 2.2 respectively. Here, we both strengthen earlier related results, and seek to make the underlying arguments more transparent. These "master theorems", which can be viewed as asymptotic results on Gaussian random matrices,

Introduction

can be adapted to analyse variants of the original AMP recursion that are geared towards more statistical problems. In this aspect, we focus on two canonical statistical settings, namely estimation of low-rank matrices in Section 3, and estimation in generalised linear models (GLMs) in Section 4. The former encompasses Sparse Principal Component Analysis (Deshpande and Montanari, 2014; Gataric et al., 2020; Jolliffe et al., 2003; Wang et al., 2016; Zou et al., 2006), submatrix detection (Ma and Wu, 2015), hidden clique detection (Alon et al., 1998; Deshpande and Montanari, 2015), spectral clustering (von Luxburg, 2007), matrix completion (Candès and Recht, 2009; Zhu et al., 2019), topic modelling (Blei et al., 2003) and collaborative filtering (Su and Khoshgoftaar, 2009). The latter provides a holistic approach to studying a suite of popular modern statistical methods, including penalised M-estimators such as the Lasso (Tibshirani, 1996) and SLOPE (Bogdan et al., 2015), as well as more traditional techniques such as logistic regression. A novel aspect of our presentation in Section 4 is that we formalise the connection between AMP and a broad class of convex optimisation problems, and then show how to systematically derive exact expressions for the asymptotic risk of estimators in GLMs. We expect that our general recipe can be applied to a wider class of GLMs than have been studied in the AMP literature to date.

To preview the statistical content in this tutorial and highlight some recurring themes, we now discuss two prototypical applications of AMP that form the basis of Sections 3 and 4 respectively. First, suppose that we wish to estimate an unknown signal $v \in \mathbb{R}^n$ based on an observation

$$A = \frac{\lambda}{n} v v^{\top} + W,$$

where $\lambda > 0$ is fixed and $W \in \mathbb{R}^{n \times n}$ is a symmetric Gaussian noise matrix. In this so-called spiked Wigner model (see Section 3.1 and the references therein), a popular and well-studied estimator of v is the leading eigenvector $\hat{\varphi}$ of A, which can be approximated via the power method, with iterates

$$v^{k+1} = \frac{Av^k}{\|Av^k\|}.$$

$$\hat{v}^k = g_k(v^k), \qquad v^{k+1} = A\hat{v}^k - b_k\hat{v}^{k-1}$$

for $k \in \mathbb{N}_0$, where we emphasise the following two characteristic features:

- (i) Each "denoising" function $g_k \colon \mathbb{R} \to \mathbb{R}$ is applied componentwise to vectors, and can be chosen appropriately to exploit different types of prior information about the structure of v (e.g., to encourage \hat{v}^k to be sparse).
- (ii) In the "memory" term $-b_k \hat{v}^{k-1}$, which is called an "Onsager" correction in the AMP literature (e.g., Bayati and Montanari, 2011; Donoho *et al.*, 2009), the scalar b_k is defined as a specific function of v^k to ensure that the iterates v^{k+1} have desirable statistical properties; see (3.3).

One way to incorporate additional structural information on v into the spiked model is to assume that its entries are drawn independently from some prior distribution π on \mathbb{R} ; for example, we can enforce sparsity through priors that place strictly positive mass at 0. Then under appropriate conditions, AMP theory guarantees that, for each k, the components of the estimate \hat{v}^k have approximately the same empirical distribution as those of $g_k(\mu_k v + \sigma_k \xi)$; here, $\xi \sim N_n(0, I_n)$ is a "noise" vector that is independent of the signal $v \in \mathbb{R}^n$, and the "signal" and "noise" parameters $\mu_k \in \mathbb{R}, \sigma_k > 0$ are determined by a scalar state evolution recursion that depends on (g_k) and the prior distribution π ; see (3.6). This distributional characterisation effectively reduces the analysis of the high-dimensional \hat{v}^k to a much simpler univariate denoising problem, where the aim is to reconstruct $V \sim \pi$ based on a single corrupted observation of the form $\mu_k V + \sigma_k G$ with $G \sim N(0,1)$ representing independent Gaussian noise. The functions g_k can then be chosen in such a way that the "effective signal-to-noise ratios" $(\mu_k/\sigma_k)^2$ are large and $g_k(\mu_k V + \sigma_k G)$ accurately estimates V. This ensures that the resulting AMP estimates $\hat{v}^k = q_k(v^k)$ have low asymptotic estimation error as $n \to \infty$.



Figure 1.1: Asymptotic mean-squared error plots for estimation of a signal $v \in \mathbb{R}^n$ with i.i.d. $U\{-1, 1\}$ entries in the rank-one spiked model, based on an AMP algorithm with denoising functions $g_k : x \mapsto \tanh(\mu_k x/\sigma_k^2)$ and spectral initialisation $(v^0 = \hat{\varphi}$ and $\hat{v}^{-1} = \lambda^{-1}\hat{\varphi}$ with $\|\hat{\varphi}\| = \sqrt{n\lambda^2(\lambda^2 - 1)_+}$. See Sections 3.2–3.3, where we also discuss how to consistently estimate λ when it is unknown (Remark 3.12). *Left*: Plot of $AMSE_k(\lambda) := \lim_{n\to\infty} \|\hat{v}^k - v\|^2/n$ against the iteration number k for the AMP estimates $\hat{v}^k \equiv \hat{v}^k_\lambda(n)$, when $\lambda = 1.7$. $AMSE_k(\lambda)$ decreases monotonically to some $AMSE_{\infty}(\lambda)$ as $k \to \infty$; see Theorem 3.10(c). *Right*: Plots of $AMSE_{-1}(\lambda) =$ $1 \wedge \lambda^{-2}$ for the pilot spectral estimator \hat{v}^{-1} and $AMSE_{\infty}(\lambda)$ for AMP, with $\lambda \in [0, 3]$. The spectral estimator undergoes the so-called BBP phase transition at $\lambda = 1$; see Section 3.1.

For instance, suppose that the entries of v are drawn uniformly at random from $\{-1, 1\}$. Then provided we initialise the AMP algorithm with $v^0 = \hat{\varphi}$ and $\hat{v}^{-1} = \lambda^{-1}\hat{\varphi}$, where $\|\hat{\varphi}\| = \sqrt{n\lambda^2(\lambda^2 - 1)_+}$, it turns out that the asymptotic mean squared error (MSE) of \hat{v}^k is minimised by choosing g_k to be the function $x \mapsto \tanh(\mu_k x / \sigma_k^2)$; see Section 3.3. Figure 1.1 illustrates that the limiting MSE of the AMP estimates \hat{v}^k decreases with the iteration number k, and in particular that they improve on the pilot spectral estimator \hat{v}^{-1} (which is agnostic to the structure of v).

As a second example, consider the linear model $y = X\beta + \varepsilon$, where $\beta \in \mathbb{R}^p$ is the target of inference, $\varepsilon \in \mathbb{R}^n$ is a noise vector, and $X \in \mathbb{R}^{n \times p}$ is a random design matrix with independent N(0, 1/n) entries. In highdimensional regimes where p is comparable in magnitude to, or even much larger than n, a popular (sparse) estimator is the Lasso (Tibshirani, 1996), which for $\lambda > 0$ is defined by

$$\hat{\beta}^{\mathrm{L},\lambda} \in \operatorname*{argmin}_{\tilde{\beta} \in \mathbb{R}^p} \bigg\{ \frac{1}{2} \|y - X\tilde{\beta}\|^2 + \lambda \|\tilde{\beta}\|_1 \bigg\}.$$

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In the literature on high-dimensional estimation, upper bounds on the prediction and estimation error of the Lasso have been obtained under suitable conditions on the design matrix X, such as the restricted isometry property or compatibility conditions (e.g., Bühlmann and van de Geer, 2011). AMP offers complementary guarantees by providing exact formulae for the asymptotic risk in the "large system limit" where $n, p \to \infty$ with $n/p \to \delta \in (0, \infty)$, and with the components of β drawn independently from a prior distribution on \mathbb{R} . To motivate the form of the AMP algorithm in this setting, first consider the iterative soft thresholding algorithm (ISTA) for solving the Lasso optimisation problem, whose update steps can be written as

$$\hat{r}^k = y - X\hat{\beta}^k, \qquad \hat{\beta}^{k+1} = \operatorname{ST}_{\lambda\eta_k}(\hat{\beta}^k + \eta_k X^{\top} \hat{r}^k) \qquad \text{for } k \in \mathbb{N}_0; (1.1)$$

here, \hat{r}^k is the current residual, $\eta_k > 0$ is a deterministic step size, and for t > 0, the soft-thresholding function $ST_t: w \mapsto sgn(w)(|w| - t)_+$ is applied componentwise to vectors. This is an instance of the generalpurpose proximal gradient method (Parikh and Boyd, 2013, Sections 4.2 and 4.3). An "accelerated" version of (1.1) called FISTA (Beck and Teboulle, 2009) bears a closer resemblance to an AMP algorithm, where the iterates of the latter are given by

$$\hat{r}^{k} = y - X\hat{\beta}^{k} + \frac{\|\hat{\beta}^{k}\|_{0}}{n}\,\hat{r}^{k-1}, \quad \hat{\beta}^{k+1} = \mathrm{ST}_{t_{k+1}}(\hat{\beta}^{k} + X^{\top}\hat{r}^{k}) \quad \text{for } k \in \mathbb{N}_{0}.$$
(1.2)

Here, each $t_k > 0$ is a deterministic threshold and $\|\hat{\beta}^k\|_0$ denotes the number of non-zero entries of $\hat{\beta}^k \in \mathbb{R}^p$. By comparison with (1.1), we observe that \hat{r}^k in (1.2) is a corrected residual, whose definition includes an additional memory term that is crucial for ensuring that the empirical distribution of the iterates can be characterised exactly. Indeed, for each fixed $k \in \mathbb{N}$, the entries of the AMP estimate $\hat{\beta}^k$ of β have approximately the same empirical distribution as those of $\operatorname{ST}_{t_k}(\beta + \sigma_k \xi)$ when p is large; here $\xi \sim N_p(0, I_p)$ is a noise vector that is independent of β , the noise level $\sigma_k > 0$ is determined by the state evolution recursion defined in (4.41) below, and the scalar denoising function ST_{t_k} induces sparsity.

Bayati and Montanari (2012) proved that in the asymptotic regime above, the AMP iterates $(\hat{r}^k, \hat{\beta}^k)$ converge in a suitable sense to a fixed point $(\hat{r}^*, \hat{\beta}^*)$, and a key property of (1.2) is that for any such fixed point,

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 $\hat{\beta}^*$ is a Lasso solution; see (4.42) below. It follows that the performance of the Lasso is precisely characterised by a fixed point of the state evolution recursion (4.41); see Theorem 4.5. Since the above properties are proved under a Gaussian design, the main utility of AMP in this setting is not so much as an efficient Lasso computational algorithm, but rather as a device for gaining insight into the statistical properties of the estimator. In Section 4, the above theory is developed as part of an overarching AMP framework for linear models and generalised linear models (GLMs).

Note that in both of the examples above, the limiting empirical distributions of the entries of the AMP iterates can be decomposed into independent "signal" and "noise" components, and the effective signal strength and noise level are determined by a state evolution recursion. In Sections 3 and 4, we show how to derive these asymptotic guarantees by applying the master theorems in Section 2 to suitable abstract recursions, which track the evolution of the asymptotically Gaussian "noise" components of the AMP iterates. We discuss various extensions in Section 5, and provide proofs in the Appendix (Section A), with supplementary mathematical background deferred to Section B. As a guide to the reader, we remark that rigorous formulations of the results in this monograph require a number of technical conditions. While we take care to state these precisely, and discuss them at appropriate places, we emphasise that these should generally be regarded as mild. We therefore recommend that the reader initially focuses on the main conclusions of the results.

The statistical roots of AMP lie in compressed sensing (Donoho *et al.*, 2009, 2013). A reader approaching the subject from this perspective can consult Montanari (2012), Tramel *et al.* (2014) and Schniter (2020) for accessible expositions of the motivating ideas and the connections with message passing algorithms on dense graphs. Alternatively, for comprehensive reviews of AMP from a statistical physics perspective, see Zdeborová and Krzakala (2016), Krzakala *et al.* (2012) and Lesieur *et al.* (2017).

In spin glass theory, an AMP algorithm was proposed as an iterative scheme for solving the Thouless–Anderson–Palmer (TAP) equations

1.1. Notation and Preliminaries

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corresponding to a Sherrington-Kirkpatrick model with specific parameters (Bolthausen, 2014; Mézard and Montanari, 2009; Mézard *et al.*, 1987; Talagrand, 2011). The estimation problem here is equivalent to one of reconstructing a symmetric rank-one matrix in a Gaussian spiked model. Bolthausen (2014) proved a rigorous state evolution result for AMP in this specific setting, by introducing a conditioning argument that became an essential ingredient in subsequent analyses of AMP (Bayati and Montanari, 2011; Berthier *et al.*, 2020; Javanmard and Montanari, 2013; Fan, 2022). See Section A.2 for a detailed discussion of this proof technique.

In this tutorial, we restrict our focus to AMP recursions in which the random matrices are Gaussian. However, as we discuss in Section 5, several recent works have extended AMP and its state evolution recursion to more general non-Gaussian settings. For matrices with independent sub-Gaussian entries, results on the "universality" of AMP were first established by Bayati *et al.* (2015) and later in greater generality by Chen and Lam (2021). In addition, to accommodate the class of rotationally invariant random matrices, a number of extensions of the original AMP framework have recently been proposed, including Orthogonal AMP (Ma and Ping, 2017; Takeuchi, 2020) and Vector AMP (Schniter *et al.*, 2016; Rangan *et al.*, 2019b), as well as the general iterative schemes of Opper *et al.* (2016), Çakmak and Opper (2019) and Fan (2022). Some of these are closely related to expectation propagation (Opper and Winther, 2005; Kabashima and Vehkaperä, 2014). In all of the above variants of AMP, the recursion is tailored to the spectrum of the random matrix.

1.1 Notation and Preliminaries

Here, we introduce some notation used throughout this tutorial, and present basic properties of Wasserstein distances, pseudo-Lipschitz functions, as well as the complete convergence of random sequences.

General notation: For $n \in \mathbb{N}$, let e_1, \ldots, e_n be the standard basis vectors in \mathbb{R}^n . For $r \in [1, \infty]$, we write $||x||_r$ for the ℓ_r norm of $x \equiv (x_1, \ldots, x_n) \in \mathbb{R}^n$, so that $||x||_r = (\sum_{i=1}^n |x_i|^r)^{1/r}$ when $r \in [1, \infty)$ and $||x||_{\infty} = \max_{1 \leq i \leq n} |x_i|$. We also define $||x||_{n,r} := n^{-1/r} ||x||_r = (n^{-1} \sum_{i=1}^n |x_i|^r)^{1/r}$ for $r \in (1, \infty)$. Let $\langle \cdot, \cdot \rangle$ and $||\cdot|| := ||\cdot||_2$ be the

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standard Euclidean inner product and norm on \mathbb{R}^n respectively, and define $\langle \cdot, \cdot \rangle_n$ to be the scaled Euclidean inner product on \mathbb{R}^n given by $\langle x, y \rangle_n := n^{-1} \langle x, y \rangle$ for $x, y \in \mathbb{R}^n$, which induces the norm $\|\cdot\|_n := \|\cdot\|_{n,2}$. We denote by $\mathbf{1}_n := (1, \ldots, 1) \in \mathbb{R}^n$ the all-ones vector and write $\langle x \rangle_n := \langle x, \mathbf{1}_n \rangle_n = n^{-1} \sum_{i=1}^n x_i$ for each $x \in \mathbb{R}^n$.

For $D \in \mathbb{N}$ and $x^1, \ldots, x^D \in \mathbb{R}^n$, we denote by $\nu_n(x^1, \ldots, x^D) := n^{-1} \sum_{i=1}^n \delta_{(x_i^1, \ldots, x_i^D)}$ the joint empirical distribution of their components, and for a function $f : \mathbb{R}^D \to \mathbb{R}$, write $f(x^1, \ldots, x^D) := (f(x_i^1, \ldots, x_i^D))$: $1 \leq i \leq n \in \mathbb{R}^n$ for the row-wise application of f to $(x^1 \cdots x^D)$.

By a Euclidean space $(E, \|\cdot\|_E)$ we mean a finite-dimensional inner product space over \mathbb{R} , equipped with the norm induced by its inner product; examples include $(\mathbb{R}^n, \|\cdot\|)$ for $n \in \mathbb{N}$ and $(\mathbb{R}^{k \times \ell}, \|\cdot\|_F)$ for $k, \ell \in \mathbb{N}$, where $\|\cdot\|_F$ is the Frobenius norm induced by the trace inner product $(A, B) \mapsto \operatorname{tr}(A^{\top}B)$.

Gaussian orthogonal ensemble: We write $W \sim \text{GOE}(n)$ if $W = (W_{ij})_{1 \leq i,j \leq n}$ takes values in the space of all symmetric $n \times n$ matrices, and has the property that $(W_{ij})_{1 \leq i \leq j \leq n}$ are independent, with $W_{ij} \sim N(0, 1/n)$ for $1 \leq i < j \leq n$ and $W_{ii} \sim N(0, 2/n)$ for $i = 1, \ldots, n$. Writing \mathbb{O}_n for the set of all $n \times n$ orthogonal matrices, we note the orthogonal invariance property of the GOE(n) distribution: if $Q \in \mathbb{O}_n$ and $W \sim \text{GOE}(n)$, then $Q^{\top}WQ \sim \text{GOE}(n)$.

Complete convergence of random sequences: The asymptotic results below are formulated in terms of the notion of *complete convergence* (e.g., Hsu and Robbins, 1947; Serfling, 1980, Chapter 1.3). This is a stronger mode of stochastic convergence than almost sure convergence, and is denoted throughout using the symbol $\stackrel{c}{\rightarrow}$. In Definition 1.1 and Proposition 1.2 below, we give two equivalent characterisations of complete convergence and introduce some associated stochastic O symbols.

Definition 1.1. Let (X_n) be a sequence of random elements taking values in a Euclidean space $(E, \|\cdot\|_E)$. We say that X_n converges completely to a deterministic limit $x \in E$, and write $X_n \xrightarrow{c} x$ or c-lim $_{n\to\infty}X_n = x$, if $Y_n \to x$ almost surely for any sequence of *E*-valued random elements (Y_n) with $Y_n \stackrel{d}{=} X_n$ for all n.

1.1. Notation and Preliminaries

We write $X_n = o_c(1)$ if $X_n \stackrel{c}{\to} 0$, and write $X_n = O_c(1)$ if $Y_n = O_{a.s.}(1)$ (i.e., $\limsup_{n\to\infty} ||Y_n||_E < \infty$ almost surely) for any sequence of *E*-valued random elements (Y_n) with $Y_n \stackrel{d}{=} X_n$ for all *n*.

Proposition 1.2. For a sequence (X_n) of random elements taking values in a Euclidean space $(E, \|\cdot\|_E)$, we have

- (a) $X_n = o_c(1)$ if and only if $\sum_n \mathbb{P}(||X_n||_E > \varepsilon) < \infty$ for all $\varepsilon > 0$;
- (b) $X_n = O_c(1)$ if and only if there exists C > 0 such that $\sum_n \mathbb{P}(||X_n||_E > C) < \infty$.

For a deterministic $x \in E$, we see that $X_n \xrightarrow{c} x$ if and only if $\sum_n \mathbb{P}(||X_n - x||_E > \varepsilon) < \infty$ for all $\varepsilon > 0$. Moreover, if $X_n \xrightarrow{c} x$, then $X_n = O_c(1)$. The proof of Proposition 1.2, along with various other properties of complete convergence and a calculus for $o_c(1)$ and $O_c(1)$ notation, is given in Section B.1; see also Remark A.1.

Wasserstein distances and pseudo-Lipschitz functions: For $D \in \mathbb{N}$ and $r \in [1, \infty)$, we write $\mathcal{P}(r) \equiv \mathcal{P}_D(r)$ for the set of all Borel probability measures P on \mathbb{R}^D with $\int_{\mathbb{R}^D} ||x||^r dP(x) < \infty$. For $P, Q \in \mathcal{P}_D(r)$, the *r*-Wasserstein distance between P and Q is defined by

$$d_r(P,Q) := \inf_{(X,Y)} \mathbb{E}(||X-Y||^r)^{1/r},$$

where the infimum is taken over all pairs of random vectors (X, Y)defined on a common probability space with $X \sim P$ and $Y \sim Q$. For $P, P_1, P_2, \ldots \in \mathcal{P}_D(r)$, we have $d_r(P_n, P) \to 0$ if and only if both $\int_{\mathbb{R}^D} ||x||^r dP_n(x) \to \int_{\mathbb{R}^D} ||x||^r dP(x)$ and $P_n \to P$ weakly (e.g., Villani, 2003, Theorem 7.12). Furthermore, for L > 0, we write $PL_D(r, L)$ for the set of functions $\psi \colon \mathbb{R}^D \to \mathbb{R}$ such that

$$|\psi(x) - \psi(y)| \le L ||x - y|| \left(1 + ||x||^{r-1} + ||y||^{r-1}\right)$$
(1.3)

for all $x, y \in \mathbb{R}^D$, and denote by $\operatorname{PL}_D(r) := \bigcup_{L>0} \operatorname{PL}_D(r, L)$ the class of pseudo-Lipschitz functions $f : \mathbb{R}^D \to \mathbb{R}$ of order r. Note that $\operatorname{PL}_D(1, L)$ is precisely the class of all (3L)-Lipschitz functions on \mathbb{R}^D , and that $\operatorname{PL}_D(s) \subseteq \operatorname{PL}_D(r)$ for any $1 \leq s \leq r$. Moreover, for any probability measure $P \in \mathcal{P}_D(r)$, we have $|\int_{\mathbb{R}^d} \psi \, dP| \leq L \int_{\mathbb{R}^D} (||x|| + ||x||^r) \, dP(x) + C$

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 $|\psi(0)| < \infty$ for all $\psi \in \operatorname{PL}_D(r, L)$. Now for $P, Q \in \mathcal{P}_D(r)$, we define

$$\widetilde{d}_r(P,Q) := \sup_{\psi \in \mathrm{PL}_D(r,1)} \left| \int_{\mathbb{R}^D} \psi \, dP - \int_{\mathbb{R}^D} \psi \, dQ \right|. \tag{1.4}$$

In Section B.4, we show (among other things) that \tilde{d}_r, d_r are metrics on $\mathcal{P}_D(r)$ that induce the same topology (Remark B.18).

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